## RMSC5102 <br> Midterm Review

Leung Man Fung, Heman
Spring, 2021

## Agenda

- Review
- Basic Knowledge
- The Black-Scholes World
- Monte Carlo Method
- Random Variable Generation
- Variance Reduction Technique
- Q\&A

Basic Knowledge

## Geometric Distribution

- $X \sim \operatorname{Geo}(p)$
- number of Bernoulli trials $X$ needed to get 1 success
- each trial has a probability of success $p$
- Pmf: $\mathbb{P}(X=x)=(1-p)^{x-1} p$ for $x=1,2, \ldots$
- Mean: $\mathbb{E}(X)=\frac{1}{p}$
- Variance: $\operatorname{Var}(X)=\frac{1-p}{p^{2}}$
- Related to rejection sampling


## Exponential Distribution

- $X \sim \operatorname{Exp}(\lambda)$
- Continuous analogue of the geometric distribution
- We adopt the rate parametrization instead of scale
- Pdf: $f(x)=\lambda e^{-\lambda x}$ for $x \geq 0$, otherwise 0
- Cdf: $F(x)=1-e^{-\lambda x}$ for $x \geq 0$
- Mean: $\mathbb{E}(X)=\frac{1}{\lambda}$
- Variance: $\operatorname{Var}(X)=\frac{1}{\lambda^{2}}$
- Memoryless property: $\mathbb{P}(X>s+t \mid X>s)=\mathbb{P}(X>t)$ for $s, t \geq 0$
- Exponential distribution is the only continuous distribution that has this property
- Useful representation: if $Y \sim \operatorname{Exp}(1)$, then $\frac{Y}{\lambda} \sim \operatorname{Exp}(\lambda)$


## Some Properties of Expectation, Variance and Covariance

- Law of the unconscious statistician: $\mathbb{E}[g(X)]=\sum_{i=1}^{n} g\left(x_{i}\right) \mathbb{P}\left(X=x_{i}\right)$
- Also holds for continuous random variables
- Translation/rescale:
- $\mathbb{E}(a X+b)=a \mathbb{E}(X)+b$
- $\operatorname{Var}(a X+b)=a^{2} \operatorname{Var}(X)$
- $\operatorname{Cov}(a X+b, c Y+d)=a c \operatorname{Cov}(X, Y), \operatorname{Cov}(X, X)=\operatorname{Var}(X)$
- Linearity of expectation: $\mathbb{E}\left(\sum_{i=1}^{n} X_{i}\right)=\sum_{i=1}^{n} \mathbb{E}\left(X_{i}\right)$
- Alternative formula for variance: $\operatorname{Var}(X)=\mathbb{E}\left(X^{2}\right)-[\mathbb{E}(X)]^{2}$
- Example: HW1 5, HW2 1


## Geometric Brownian Motion

- SDE: $\mathrm{dS}_{\mathrm{t}}=r S_{t} d t+\sigma S_{t} d W_{t} \Rightarrow \mathrm{~S}_{\mathrm{t}}=S_{0} e^{\left(r-\frac{1}{2} \sigma^{2}\right) t+\sigma W_{t}}$
- Use $S_{T}=S_{0} e^{\left(r-\frac{1}{2} \sigma^{2}\right)^{T+\sigma \sqrt{T} Z}}$ in simulation to avoid simulating intermediate prices
- Algorithm:
- Generate $Z \sim N(0,1)$
- Set $S_{T}=S_{0} e^{\left(r-\frac{1}{2} \sigma^{2}\right) T+\sigma \sqrt{T} Z}$
- Example: HW1 1, 2; HW2 5
- Question may specify stock price dynamic other than GBM
- Use the given dynamic to simulate $S_{T}$ like generating random variables


## The BlackScholes World

## Risk Neutral Valuation

- $V_{t}=e^{-r(T-t)} \mathbb{E}\left[f\left(S_{t}, t\right)\right]$
- Take expectation w.r.t. real world probability?
- E.g., with insider info you know price of a certain stock will likely go up
- Problem of the discount rate
- If real world probability is used, discount rate has to accommodate the level of risk (think about the discount rate you use in DCF)
- If risk neutral probability is used, discount rate = risk free rate (observable)
- Just give you another way of looking at risk neutral approach
- The above formula always hold. Why do we need model like GBM then?
- Because $r, t, T, f(\cdot, \cdot)$ are known/observable but $S_{t}$ is not


## Black-Scholes-Merton Model

- Black-Scholes formula
- $C\left(S_{t}, t\right)=\Phi\left(d_{1}\right) S_{t}-\Phi\left(d_{2}\right) K e^{-r(T-t)}$,
- $P\left(S_{t}, t\right)=K e^{-r(T-t)}-\mathrm{S}_{\mathrm{t}}+C\left(S_{t}, t\right)=\Phi\left(-d_{2}\right) K e^{-r(T-t)}-\Phi\left(-d_{1}\right) S_{t}$
- $d_{1}=\frac{1}{\sigma \sqrt{T-t}}\left[\ln \left(\frac{S_{t}}{K}\right)+\left(r+\frac{\sigma^{2}}{2}\right)(T-t)\right]$
- $d_{2}=d_{1}-\sigma \sqrt{T-t}$
- Note that $P\left(S_{t}, t\right)$ is derived from put-call parity
- Put-call parity: $C_{E}-P_{E}=S-K e^{-r(T-t)}$
- Implied volatility
- Value of volatility when back-solving an option pricing model (such as BS) with current market price


## Monte Carlo Method

## Key Idea

Use repeated random sampling to obtain numerical estimate

The estimate is usually average in our course
Example: estimate $\pi$ (picture credit: nicoguaro)



## Standard Monte Carlo

- HW2 5a: price a European call option
- Recall payoff function is $\max \left(S_{T}-K, 0\right)$
- Estimate $\mathbb{E}\left[\max \left(S_{T}-K, 0\right)\right]$ by sample average $\frac{1}{n} \sum_{i=1}^{n} \max \left(S_{T}^{(i)}-K, 0\right)$
- Algorithm
- 1) Generate $Z \sim N(0,1)$
- 2) Set $S_{T}=S_{0} e^{\left(r-\frac{1}{2} \sigma^{2}\right)^{T+\sigma \sqrt{T}} Z}$
- 3) Compute $\boldsymbol{\pi}_{i}=\max \left(\boldsymbol{S}_{\boldsymbol{T}}-\boldsymbol{K}, \mathbf{0}\right)$
- 4) Repeat 1 to 3 for $i=1, \ldots, n$
- 5) Option price $=\frac{\mathrm{e}^{-\mathrm{rT}}}{\mathrm{n}} \sum_{i=1}^{n} \pi_{i}$


## Stopping a Simulation

- Margin of error: terminate a simulation when $\frac{s^{2}}{\sqrt{n}} \leq d$
- Where $s^{2}$ is the sample variance and $d$ is the maximum tolerable error
- This is essentially based on the Central Limit Theorem
- Law of large numbers (WLLN): Let $X_{1}, \ldots, X_{n}$ be i.i.d. random variables with mean $\theta$ and variance $\sigma^{2}$, then $\bar{X}_{n} \approx \theta$ as $n \rightarrow \infty$
- Central limit theorem: Let $X_{1}, \ldots, X_{n}$ be i.i.d. random variables with mean $\theta$ and finite variance $\sigma^{2}$, then $\bar{X}_{n} \approx \mathrm{~N}\left(\theta, \frac{\sigma^{2}}{n}\right)$ as $n \rightarrow \infty$


## Random

 Variable Generation
## Key ideas

- Monte Carlo methods rely on repeated random sampling
- Need way(s) to generate different random variables
- Only Unif $(0,1)$ and $N(0,1)$ can be generated without any algorithm
- Technically only Unif( 0,1 ) (pseudorandom number)
- $\mathrm{N}(0,1)$ is generated by special algorithm as well, e.g., Box-Muller transformation
- They are omitted in this course
- Modulo operation: find the remainder of a division; denoted by mod.
- E.g., $97=7(\bmod 10), 25=1(\bmod 8)$
- Use inverse transform/rejection sampling if other distribution appears


## Inverse Transform

- Probability integral transform: if $X$ is a continuous random variable with cdf $F_{X}$, then $Y=F_{X}(X) \sim \operatorname{Unif}(0,1)$
- Therefore, if we know $X \sim F_{X}$ (i.e., the cdf), we can generate $X$ out of $U \sim \operatorname{Unif}(0,1)$
- Need to derive cdf if only pdf is given
- Algorithm (discrete)
- Generate $U \sim \operatorname{Unif}(0,1)$
- $X=x_{j}$ if $\sum_{i=0}^{j-1} p_{i} \leq U<\sum_{i=0}^{j} p_{i}$
- Algorithm (continuous)
- Generate $U \sim \operatorname{Unif}(0,1)$
- $X=F_{X}^{-1}(U)$ assuming the inverse exists
- Example: HW2 1c, 4a


## Rejection Sampling

- If we can simulate $Y \sim G_{Y}$ easily, we can use the proportional distribution as a basis to simulate $X$ with pdf $f(x)$
- Algorithm
- 1) Find $c=\max _{y} \frac{f(y)}{g(y)}$
- 2) Generate $Y_{i}$ from a density g: $U_{1} \sim \operatorname{Unif}(0,1) \Rightarrow Y_{i}=G^{-1}\left(U_{1}\right)$
- 3) Generate $U_{2} \sim \operatorname{Unif}(0,1)$
- 4) If $\mathrm{U}_{2} \leq \frac{1}{c} \cdot \frac{f\left(Y_{i}\right)}{g\left(Y_{i}\right)^{\prime}}$ set $X_{i}=Y_{i}$, otherwise return to 2
- Example: HW2 3, 4b
- Inverse transform is rejection sampling with $c=1$
- Because inverse transform simulate from $F$ directly (always accept)


## Variance Reduction Technique

## Antithetic Variables

- If we are able to generate negatively correlated underlying random variables, the estimator can have lower variance as compared with independent samples
- This requires the target function $h(x)$ to be monotone
- Show $h^{\prime}(x) \geq 0$ or $h^{\prime}(x) \leq 0$ within the target range for monotonicity
- As $h(x)$ is monotone, $\operatorname{Cov}[h(U), h(1-U)] \leq 0$ where $U \sim \operatorname{Unif}(0,1)$
- As half of your variables are antithetic, you only need to generate $\frac{n}{2}$ numbers for $n$ samples


## Antithetic Variables

- Algorithm:
- 1) Generate $U \sim \operatorname{Unif}(0,1)$
- 2) Set $X_{i}=F^{-1}(U), Y_{i}=F^{-1}(1-U)$ (note: want $X, Y$ same distribution but negative correlation)
- 3) Repeat 1 and 2 for $n$ times
- 4) $\hat{\theta}=\frac{1}{2 n} \sum_{i=1}^{n}\left[h\left(X_{i}\right)+h\left(Y_{i}\right)\right]$
- Note:
- $F^{-1}(U)$ is monotone in general as caf is monotone
- Hence $h\left[F^{-1}(U)\right]$ is monotone if $h(\cdot)$ is monotone


